Data Augmentation MCMC for Bayesian Inference from Privatize Data¹

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Workshop on Differential Privacy and Statistical Data Analysis
Fields Institute

¹Ju, Awan, G., & Rao. (2022). ArXiv:2206.00710.

Collaborators



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Jordan A. Awan Purdue University

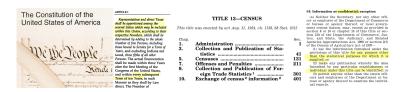


Vinayak Rao Purdue University

Privacy: a challenge in modern data curation

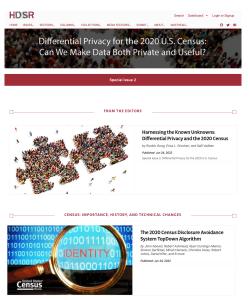
Modern data curators seek to meet two goals at once:

- 1. To **disclose** key statistics/use cases of the database, in accordance with its legal, policy, and/or ethical mandates.
- 2. To protect the **privacy** of individuals with trust-worthy guarantees.

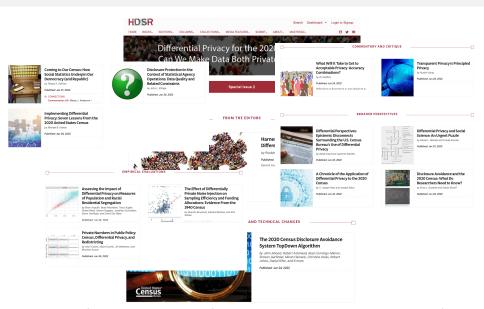


For example, the U.S. Census Bureau bears the constitutional mandate to enumerate the population every 10 years for apportionment. It is also bound by Title 13 of U.S. Code to protect respondent confidentiality.

The U.S. Census Bureau adopts differential privacy



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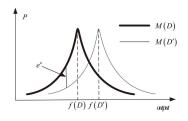


Harvard Data Science Review (https://hdsr.mitpress.mit.edu)

The mechanism of differential privacy

A random function $s_{dp}(x, r)$ is said to be ϵ -differentially private² if for all *neighboring* databases (x, x') and all possible state a,

$$\frac{\Pr\left(s_{dp}\left(\boldsymbol{x},\boldsymbol{r}\right)=a\mid\boldsymbol{x}\right)}{\Pr\left(s_{dp}\left(\boldsymbol{x}',\boldsymbol{r}\right)=a\mid\boldsymbol{x}'\right)}\leq\exp\left(\epsilon\right).$$



https://www.ons.gov.uk/peoplepopulationandcommunity

That is, differentially private mechanisms conceal the confidential data x by infusing crafted noise r into the data product s_{dp} for release:

$$\mathbf{x} \longrightarrow s_{\mathrm{dp}} \left(\mathbf{x}, \mathbf{r} \right)$$

²Dwork et al. (2006). Calibrating noise to sensitivity in private data analysis. TCC (pp 265-284)

Differential privacy: benefits and challenges

- ✓ Provability: differential privacy guarantees are formal and verifiable;
- √ Transparency: The probabilistic specification of the privacy mechanism can publicized without sabotaging the privacy guarantee.
- Statistical implication: transparency is necessary for drawing principled inference from privatized data.³

How do we leverage the privacy mechanism for statistical inference?

³G. (2022). Transparent Privacy is Principled Privacy. HDSR, Special Issue 2.

Situating our (statistical \times privacy) framework

x is the **truth**

$$s_{\rm dp} \mid x \sim \eta(s_{\rm dp} \mid x)$$

- ► Infer *x* based on *s*_{dp};
- η is the only source of uncertainty.

x is a sample

$$x \mid \theta \sim f(x \mid \theta)$$
$$s_{dp} \mid x \sim \eta(s_{dp} \mid x)$$

- ▶ Infer θ based on s_{dp} .
- Uncertainty stems from η , f, and beyond

design framework

Choose the best mechanism (η) + estimator combo:⁴

$$\hat{\theta}_{\text{design}}\left(s_{\text{dp}}^{*}\left(x\right)\right)$$

adjustment framework

For a given mechanism (η) , perform the best inference:

$$\hat{\theta}_{\mathrm{adjust}}\left(s_{\mathrm{dp}}\left(x\right)\right)$$

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⁴Slavković & Seeman. (2022). Statistical Data Privacy: A Song of Privacy and Utility. ArXiv:2205.03336.

Without privacy:

- ▶ Likelihood inference: $\ell(\theta \mid x) = f(x \mid \theta)$;
- ▶ Bayesian inference: $p(\theta \mid x) \propto p(\theta)f(x \mid \theta)$.

With privacy:

► The marginal likelihood integrates over X, the entire space of confidential databases:

$$\ell(\theta \mid s_{dp}) = \int_{\mathcal{X}} \eta(s_{dp} \mid x) f(x \mid \theta) \ dx.$$

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Existing solutions

Approximations:

- Asymptotic approximation (Wang et al., 2018)
- Variational approximation (Karwa et al., 2016)
- Parametric bootstrap (Ferrando et al., 2020)

Exact but limited:

- ► Integrate exactly (Awan & Slavković, 2018, 2020)
- MCMC with latent sufficient stat (Bernstein & Sheldon, 2018, 2019)
- ► Exact inference with approximate computation (Gong, 2019)

Our Goal

An **efficient** and **user-friendly** sampler for the **exact** posterior that works for **general** choices of the data model f and the prior p.

A traditional Gibbs sampler

Problem #1: If *n* individuals each contribute *d* features, then

$$\mathcal{X} = \mathbb{X}^{n \times d}$$
.

The likelihood may be intractable, and the posterior *doubly* intractable.

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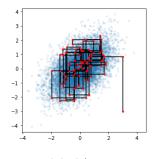
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Data Augmentation (DA). Iterate the following:

- 1: sample $\theta \mid x, s_{dp} \stackrel{d}{=} \theta \mid x$
- 2: **for** i = 1, ..., n **do**
- 3: sample $x_i \mid x_{-i}, \theta, s_{dp}$
- 4: end for



jessicastringham.net

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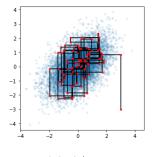
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Problem #2: the conditional dist. $x \mid \theta$, s_{dp} is both f- and η -specific.

A general Metropolis-within Gibbs sampler

Solution. Propose $x \mid \theta$ instead (or $x_i \mid \theta$ under conditional independence):

One Iteration of the privacy-aware Metropolis-within-Gibbs sampler

- 1: update $\theta \mid x$
- 2: **for** i = 1, ..., n **do**
- 3: propose $x_i^* \mid \theta$
- 4: accept x_i^* with probability

$$\alpha(\mathbf{x}_i^{\star} \mid \mathbf{x}_i, \mathbf{x}_{-i}, \theta) = \min \left\{ \frac{\eta(\mathbf{s}_{\mathrm{dp}} \mid \mathbf{x}_i^{\star}, \mathbf{x}_{-i})}{\eta(\mathbf{s}_{\mathrm{dp}} \mid \mathbf{x}_i, \mathbf{x}_{-i})}, 1 \right\}$$

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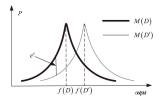
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5: end for

If η is ϵ -DP, then for all θ, x_{-i}, x_i and x_i^* :

$$\alpha(x_i^* \mid x_i, x_{-i}, \theta) \ge \exp(-\epsilon).$$

As ϵ decreases (more privacy), acceptance rate α increases (more computational efficiency)!



e.g. for
$$\epsilon = 1$$
, $\alpha \geq 36.7\%$.

Requirements, run time, and efficiency

The proposed sampler requires:

► **Assumption 1**. The analyst knows how to sample the posterior if the data aren't privatized, i.e. she has a Markov kernel targeting $p(\theta \mid x)$.

Furthermore, if we have

▶ **Assumption 2** (*Record Additivity*). The privacy mechanism can be written as $\eta(s_{dp} \mid x) = g\left(s_{dp}, \sum_{i=1}^{n} t_i(x_i, s_{dp})\right)$ for some known and tractable functions g, t_1, \ldots, t_n , then:

The Gibbs sampler requires O(n) number of operations to update the full latent database according to $p(x \mid \theta, s_{dn})$.

Note

- Even without privacy, one round of an MCMC procedure typically takes O(n) time;
- Many commonly used DP mechanisms satisfy record additivity.
 e.g. additive, exponential mechanism, objective perturbation, etc.

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Ergodicity of the proposed sampler

Under the conditions

- A1 the prior distribution is proper and $p(\theta) > 0$ for all θ in $\Theta = \{\theta \mid f_{\theta}(x) > 0 \text{ for some } x\};$
- A2 the model is such that the set $\{x: f(x\mid\theta)>0\}$ does not depend on θ ; and
- A3 the privacy mechanism satisfies $\eta(s_{\rm dp}\mid x)>0$ for all $x\in\mathbb{X}^n$, the Metropolis-within-Gibbs sampler on the joint space $(\mathbb{X}^n\times\Theta)$ is *ergodic* and it admits $p(x,\theta\mid s_{\rm dp})$ as the unique limiting distribution.

Furthermore, if one can directly sample from $p(\theta \mid x)$, then the resulting (x,θ) chain as well as the marginal chains are *geometrically ergodic* if there exists $0 < a \le b < \infty$ such that $a \le f(x \mid \theta) \le b$ for all θ and x.

Application: a naïve Bayes classifier

- $ightharpoonup x = (x_1, \dots, x_K)$ are *features*, each taking values in $\{1, \dots, \mathcal{J}_K\}$;
- \triangleright $y \in \{1, \ldots, I\}$ is the *class*;
- ▶ The non-private data consists of *n* i.i.d. copies of (x, y).
- ▶ Goal: predict the class given the features: $Pr(y \mid x)$.
- ▶ The naïve Bayes classifier assumes $Pr(x \mid y) = \prod_{k=1}^{K} Pr(x_k \mid y)$;
- ▶ Release the noisy counts: $s_{dp} = \{n_{ijk} + \text{Laplace}(2K/\epsilon)\}_{ijk}$.

Sufficient statistics of the Naive Bayes model.

An example of the parameters of the Naive Bayes model for a $2 \times 2 \times K$ table.

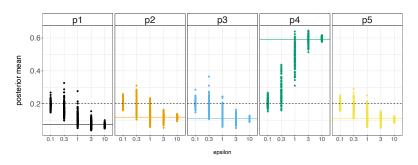
Simulation setup

For the simulation, set

- ightharpoonup N = 100 (number of samples);
- ightharpoonup I = 5 (number of classes);
- ightharpoonup K = 5 (number of features);
- ▶ \mathcal{J}_K = 3 (number of options for each feature);
- $\epsilon \in \{.1, .3, 1, 3, 10\};$
- ▶ Prior for all parameters: Dirichlet(2, ..., 2).

Posterior mean

- Fix a confidential dataset;
- \triangleright Create 100 privatized datasets at each ϵ value;
- ▶ Run one chain for each privatized dataset for 10,000 iterations;
- ▶ For each chain, calculate the posterior mean for $p_i = Pr(y = i)$.



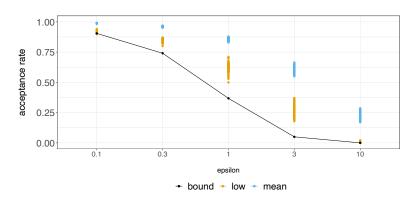
Frequentist coverage

Table. Frequentist coverage of a 90% credible interval for $p_i = \Pr(y = i)$ at different ϵ values. 100 replicates per ϵ value.

| ϵ | $p_1 = .097$ | $p_2 = .148$ | $p_3 = .145$ | $p_4 = .446$ | $p_5 = .163$ |
|------------|--------------|--------------|--------------|--------------|--------------|
| .1 | 1 | 1 | 1 | .36 | 1 |
| .3 | .97 | 1 | 1 | .59 | 1 |
| 1 | .94 | .99 | .97 | .83 | .98 |
| 3 | .95 | .91 | .97 | .89 | .93 |
| 10 | .92 | .88 | .94 | .92 | .9 |

Empirical acceptance rates

- ▶ 100 chains at each ϵ value;
- ► Each chain ran for 10,000 iterations;
- Minimum (orange) and mean (blue) acceptance rates for each chain.



Summary

An MCMC framework for Bayesian inference from privatized data:

- **Exact**: targets the correct posterior distribution;
- ► **General**: applicable to a wide range of statistical models and privacy mechanisms;
- User-friendly: mechanism independent, no (further) tuning parameters.

the privacy-efficiency alignment

Smaller ϵ implies higher acceptance rate: allowing the "free exploitation" of the privacy guarantee for computational efficiency.

Thank you

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Ju, Awan, G., & Rao. (2022). Data Augmentation MCMC for Bayesian Inference from Privatize Data.⁵ ArXiv:2206.00710.

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